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# EVIDENCE FOR OCEANIC CONTROL OF INTERANNUAL CARBON CYCLE FEEDBACKS

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Large-scale carbon-cycle feedbacks within Earth's climate system can ABSTRACT. be inferred from the statistical correlation of atmospheric CO<sub>2</sub> and other climate observations. These statistical relationships can serve as validation targets for global carbon-cycle models. Fourier-transform coherence between atmospheric CO<sub>2</sub> measured at Mauna Loa, Hawaii, and Hadley Centre global-average temperatures changed in the late 20th century at interannual frequencies, from a 6-month time lag to a  $90^\circ$ phase lag that scaled CO<sub>2</sub> fluctuations to a time-integral of the global-average temperature anomaly. Wavelet coherence estimates argue that this change occurred with a recognized ocean-circulation climate transition during the late 1970s. General features of these CO<sub>2</sub>-temperature correlations are confirmed using global-average temperature from other sources and atmospheric CO<sub>2</sub> measured at other locations, though only the Mauna Loa  $CO_2$  record is long enough to resolve well the coherence properties before the 1970s transition. The  $CO_2$ -coherence phase for the globalaverage surface-air temperature time series from NASA-GISS and the lowertroposphere temperature series from the MSU satellite is more complex than for the Hadley-Centre dataset, the only estimate that incorporates sea-surface temperature (SST) observations. Near f = 0.25 cyc/year, 4-year oscillation period, the CO<sub>2</sub>coherence is particularly strong for the Hadley-Centre gridpoint temperature-anomaly time series from low-latitude oceans. This suggests that sea-surface temperature is a primary driver of the correlation, at least for the 0.2 < f < 0.5 cyc/yr bandpass where the El-Nino/Southern-Oscillation (ENSO) climate process dominates. Outside the ENSO bandpass coherence is significant between 14 long-running GLOBALVIEW CO<sub>2</sub>-observing sites and the sea-level-pressure-based Southern Oscillation Index (SOI) and North Atlantic Oscillation (NAO) time series, consistent with wind stress and mixed-layer-thickness influences on ocean-atmosphere CO2 flux, independent of temperature fluctuations. Evidence for terrestrial biosphere influence is strongest in the leading principal component of GLOBALVIEW  $CO_2$ -variability at f = 0.25 cpy, where a larger amplitude and a 4-month phase shift distinguish the mid- and highlatitude Northern Hemisphere CO<sub>2</sub> fluctuations from those of the tropics and the Southern Hemisphere. The terrestrial signal we infer, however, coheres more strongly with oceanic-gridpoint temperatures than to continental-gridpoint temperatures.

Key words: Carbon cycle, greenhouse feedbacks, ocean circulation, climate change, El-Nino/Southern-Oscillation (ENSO), North Atlantic Oscillation (NAO) and interannual temperature fluctuations

### INTRODUCTION

Carbon-cycle feedbacks influence the absorption of anthropogenic  $CO_2$  by the Earth system. Understanding these feedbacks is key for monitoring and perhaps enhancing net carbon sequestration in the future. Uncertainties in carbon-cycle models remain large, so it can be useful to examine correlations between climate and carbon-cycle time series, independent of computer-simulation models. Data correla-

tions do not imply causality, but they can be used to support or undermine causal hypotheses, and to suggest new hypotheses. The explosive growth in  $CO_2$  data collection encourages the intensive analysis of transient climate events (for example, Chavez and others, 1999), but Earth's climate system contains processes that operate on a range of time scales. It is therefore pertinent to estimate correlation and coherence across different frequencies with the tools of spectrum analysis, to complement and contrast time-domain correlations. Many familiar causal relationships embedded within stochastic signals can be expressed via simple models of coherence and correlation amplitude and phase in the frequency domain. For illustrative examples with open-source software, see the online supplement of Park (2009).

Given the complexities of Earth's carbon cycle, it is perhaps surprising that atmospheric CO<sub>2</sub> observations correlate significantly with global-average properties of Earth's meteorology. Nevertheless, Kuo and others (1990) found strong spectral coherence between 1958 to 1988 global-average temperature fluctuations  $\Delta T_i$  and interannual CO<sub>2</sub> variability measured at Mauna Loa, Hawaii, the longest CO<sub>2</sub> time series with monthly sampling (Keeling, 2008). Revisiting the problem 20 years later, Park (2009) found a subtle, but significant, transition in the correlation. Starting from a fixed 6-month time lag during 1958 to 1988, similar to that reported by Kuo and others (1990), during 1979 to 2008 a constant 90° phase lag across frequencies 0.2 <f < 0.7 cyc/yr related CO<sub>2</sub> roughly to a simple integral of temperature fluctuations. Was this transition caused by (1) the gradual saturation of the oceanic sink for atmospheric CO<sub>2</sub> (Le Quere and others, 2007; Schuster and others, 2009; Khatiwala and others, 2009); (2) an abrupt transition of the ocean circulation (Trenberth, 1990; Graham, 1994; Park and Mann, 2000; Meehl and others, 2009); or (3) a statistical artifact related to a particular CO<sub>2</sub> observation site and a particular global-average temperature time series? This study argues for option (2), and also explores additional relationships among climate and carbon-cycle processes.

In this paper, we examine correlations among a set of earth-system data series, principally with Fourier and wavelet-based spectrum analysis, to relate interannual CO<sub>2</sub>-temperature coherence to carbon-cycle processes in the last half-century. We find that interannual ocean-climate fluctuations typically have the simplest coherence relationships with CO<sub>2</sub>, which validates the notion that they are the primary driver of interannual CO<sub>2</sub> fluctuations. Land regions either exert secondary CO<sub>2</sub> feedbacks or operate in close linkage with ocean-climate fluctuations. The Background section reviews past work on interannual correlations between climate and atmospheric CO<sub>9</sub> data. Here we note that 1990s studies confirmed the Kuo and others (1990) coherence between  $CO_2$  and temperatures, suggesting a linkage via oceanic  $CO_2$  exchange, but were followed after 2000 by many inverse studies of CO<sub>2</sub> variability that neglected oceanic processes in favor of terrestrial-biosphere influence. The Data section discusses the data time series used in this study. Fourteen globally distributed 30-year CO<sub>2</sub> time series are available, matched by global-average temperature time series from satellite measurements, from meteorological stations, and from a mix of meteorological stations and shipboard sea-surface temperature measurements. The Methods section reviews spectrum-analysis methods, based on multiple-taper spectrum analysis and a related multiple-wavelet algorithm, to correlate pairs of time series with coherence estimates, or multiple records with a spectral-domain singular-value decomposition (SVD). The section titled The Coherence Between Atmospheric  $CO_2$  and Global Temperature compares CO<sub>2</sub>-coherence with global-temperature time series from competing sources, expanding the restricted focus of Park (2009). Although globalaverage temperature times series from multiple sources show significant coherence with atmospheric CO<sub>9</sub>, the amplitude and phase relationships with the time series provided by the Climate Research Unit at the University of East Anglia lead to the

# Jeffrey Park

simplest time-domain interpretation. The section titled When and How Did the Shift in Coherence Phase Occur? uses wavelet spectra and coherence to elucidate the coherence transition reported by Park (2009), linking it temporally to the reported ocean-circulation shift in the late 1970s in the northern Pacific. The section titled The Coherence Between Atmospheric CO<sub>2</sub>, NAO and SOI addresses CO<sub>2</sub>-coherence with the Southern Oscillation Index (SOI) and North Atlantic Oscillation (NAO). We detect CO<sub>2</sub>-coherence with distinct phase lags for SOI inside and outside the El-Nino/ Southern-Oscillation (ENSO) passband, and detect correlation between NAO and the global carbon cycle at decadal time scales. The Discussion section discusses the results. The Appendix further examines the spatial variability of CO<sub>2</sub> fluctuations on ENSO time scales, illustrating the strong coherence of CO<sub>2</sub> measurements worldwide with tropical SST fluctuations.

# BACKGROUND

At interannual time scales,  $CO_2$  fluctuations lag behind global-temperature fluctuations (Kuo and others, 1990; Martin and others, 1994; Dettinger and Ghil, 1998; Adams and Piovesan, 2005; Park, 2009). Causative carbon-cycle feedback processes potentially include the response of the terrestrial biosphere to local temperature fluctuations, drought and fire (Buermann and others, 2007; Field and others, 2009; Zhao and Running, 2010), sea-surface temperatures, wind stress and ocean circulation (Chavez and others, 1999; Obata and Kitamura, 2003; Takahashi and others, 2009; Gruber and others, 2009), silicate weathering (Stallard, 1998; Gaillardet and others, 1999; Hilley and Porder, 2008; Raymond and others, 2008), volcanic eruptions (Lucht and others, 2002; Reichenau and Esser, 2003; Gu and others, 2003), and organiccarbon transport and burial in terrestrial watersheds (Stallard, 1998; Gudacz and others, 2010). Even with the advent of a global network of CO<sub>2</sub>-monitoring stations (Masarie and Tans, 1995; GLOBALVIEW-CO<sub>2</sub>, 2009), and the inspiring data-fits and visualization tools of modern data-assimilation models (Peters and others, 2007), the multiplicity of contributing factors formally makes the inverse-modeling of CO<sub>2</sub> variability an ill-posed problem (for example, Parker, 1994, Chapter 2).

The land biosphere and the ocean are estimated to sequester comparable quantities of anthropogenic carbon, 1.5 to 2 GtC/yr each (Lee and others, 1998; Battle and others, 2000; Canadell and others, 2007; Houghton, 2007). Although carbon-cycle models focused on oceanic exchanges have been applied to interannual CO<sub>2</sub> variability (for example, McKinley and others, 2006), several interpretations of interannual CO<sub>2</sub> fluctuations have focused on land-biosphere response and neglect oceanic responses (Lee and others, 1998; Bousquet and others, 2000; Russell and Wallace, 2004; Zeng and others, 2005; Buermann and others, 2007; Qian and others, 2008). One possible justification for the narrow focus is that the large annual cycle in  $CO_2$ time series is dominated by the seasonal cycle of terrestrial plant growth (Peters and others, 2007). Another justification is the prediction of weak interannual variability in the net oceanic CO<sub>2</sub> uptake in some carbon-cycle models, for example Lee and others (1998). However, carbon-cycle models do not all predict a near-constant oceanic CO<sub>2</sub> uptake (for example, Rodenback and others, 2003; Raynaud and others, 2006; Park and others, 2006), and attempts by Butler and others (2007) to fit CO<sub>2</sub> variability in the vegetation-poor boreal Southern Hemisphere led them to conclude that oceanic CO<sub>2</sub> flux was crucial. Boyce and others (2010) reported correlation between ENSO and historical phytoplankton abundance that suggests additional oceanic processes at work. The recent shift in the annual  $CO_2$  cycle (Piao and others, 2008) suggests that enhanced  $CO_9$ -uptake by the terrestrial biosphere may largely be cancelled by  $CO_9$ respiration from enhanced plant decay, reducing the contribution of terrestrial "greening" to interannual variability. Moreover, interannual CO2 fluctuations in the terrestrial biosphere are linked more closely with oceanic ENSO variability than other

factors (Nagai and others, 2007; Qian and others, 2008; Cadule and others, 2010). Ice-cover considerations (Olsen and others, 2003) and shipboard measurements (Watson and others, 2009) suggest that the important North Atlantic locus of  $CO_2$ -uptake varies year-to-year by 50 percent or more.

Attempts to verify temperature-CO<sub>2</sub> coherence in environmental time series have highlighted linkages via the tropical ocean. Interannual CO<sub>2</sub>-variations were associated with ENSO using spectral coherence (Martin and others, 1994), lagged crosscorrelation (Dettinger and Ghil, 1998), empirical-orthogonal-function (EOF) decomposition (Lintner, 2002), and other time-domain methods. Using the  $5^{\circ} \times 5^{\circ}$  Had-Crut3 gridded temperature times series (Brohan and others, 2006), Park (2009) found strong spectral coherence between interannual sea-surface temperature (SST) fluctuations and CO<sub>2</sub>-fluctuations at three widely-spaced locations (Mauna Loa, Hawaii; Barrow, Alaska; South Pole). Appendix A shows additional examples. Weak spectral coherence was found with continental temperature fluctuations. These results appear at variance with the principal causative hypotheses for interannual  $CO_2$ fluctuations considered by many global carbon-cycle modeling studies (Lee and others, 1998; Bousquet and others, 2000; Russell and Wallace, 2004; Zeng and others, 2005; Buermann and others, 2007; Qian and others, 2008). The actual disagreement may be less drastic, because other climate processes are correlated with tropical SST via ENSO.

Transition from a simple time lag to a quadrature relationship may explain why recent estimates of time-domain correlation between CO<sub>2</sub> and SST fluctuations have failed to find a link, for example, Adams and Piovesan (2005). A broadband 90° phase relationship between two time series is best detected with frequency-domain correlation, rather than time-domain correlation. Moreover, simple models for surface-ocean pCO<sub>2</sub> and CO<sub>2</sub>-exchange at the sea surface (Nevison and others, 2008; Takahashi and others, 2009; Gruber and others, 2009) imply that tropical SST fluctuations should affect the time-derivative of atmospheric  $CO_{2}$ , consistent with the  $90^{\circ}$  interannual coherence phase since 1979 reported by Park (2009)-see also Butler and others (2007). Less gas can be dissolved in warmer surface waters, so the rate of CO<sub>2</sub> uptake by the ocean, that is, its time derivative, should correlate directly with SST. A similar correlation with the time-derivative of atmospheric CO<sub>2</sub> would be expected if tropical SST fluctuations caused simultaneous fluctuations in sea-surface wind stress, phytoplankton abundance, cyclic removal of sea-ice, or other processes that influence the atmosphere ocean CO<sub>2</sub> flux. Possible relationships with the land biosphere are likely to interact with the seasonal cycle of plant growth, so a direct correlation between temperature and CO<sub>2</sub> flux cannot be posited with the same confidence. Nevertheless, we find clear support for terrestrial biosphere influence in the ENSO passband, but with timing linked to SST via ENSO teleconnections. Outside the ENSO bandbass, we find simpler correlations with the climate indices based on sea-level pressure differences (SOI, NAO) than with temperatures, as discussed in the Coherence Between Atmospheric CO<sub>2</sub>, NAO and SOI section.

### DATA

We obtained monthly data for atmospheric  $CO_2$  concentrations at Mauna Loa, Hawaii and the South Pole from the Carbon Dioxide Information Analysis Center (Keeling and others, 2009; http://cdiac.ornl.gov/trends/co2/). The Mauna Loa  $CO_2$ time series is continuous from 1958 to present, but serious gaps in South Pole data collection occur before 1965. The GLOBALVIEW data product (GLOBALVIEW-CO<sub>2</sub>, 2009; Masarie and Tans, 1995; http://www.esrl.noaa.gov/gmd/ccgg/GLOBALVIEW/, provides quasi-weekly  $CO_2$  data (48 samples/yr) for individual stations of a global network, including Mauna Loa (**mlo**) and South Pole (**spo**). GLOBALVIEW-CO<sub>2</sub> offers two multivariate time-series data products for the 1979 to 2008 interval, one in which the data series at individual observing sites are made gap-free and extended with a global interpolation scheme (EXT data product), and another in which  $CO_2$  at the marine boundary layer (MBL data product) is estimated at a set of 41 latitude values. Although  $CO_2$  measurements at a particular location often exhibit local effects (Thoning and others, 1989), at annual and interannual time scales the observations average over broad regions of both land and ocean. See Appendix A for verification of long-range correlations.

We used temperature data series from contrasting sources. The HadCrut3 global gridded monthly temperature anomaly data set (Brohan and others, 2006) is consolidated to form global- and hemisphere-average monthly temperature-anomaly data sets (http://www.cru.uea.ac.uk/cru/data/). For the gridded data we interpolated any data gaps of three months or fewer at individual gridpoints. The 1958 to 2008 Mauna Loa data series could be correlated against 1194 gridpoints, and the 1979 to 2008 GLOBALVIEW data series could be correlated against 1316 gridpoints. We restored the mean 1961 to 1990 yearly cycle to the HadCrut3 temperature data to facilitate comparison with the  $CO_2$  time series, which contain the annual cycle. We used GISTEMP temperature data from the NASA Goddard Institute for Space Science (http://data.giss.nasa.gov/gistemp/), based on historical data from meteorological stations. The GISTEMP data product provides global, hemispherical and zonal averages since 1850 (Hansen and others, 1999), as well as a code to produce temperature time series at geographical gridpoints. NASA-GISS offers data products both with, and without, the historical SST measurements compiled by the Hadley Centre. We chose the data set that does not include the SST data to decrease overlap with the Hadley-Centre data. There are two sources for lower tropospheric temperature data from the Microwave Sounding Unit (MSU) satellite, the University of Alabama, Huntsville (UAH), National Space Science and Technology Center (http://www.nsstc-.uah.edu/data/msu/) and Remote Sensing Systems (RSS, http://www.remss.com). This satellite data is available only from 1979 (Spencer and Christy, 1990; Mears and others, 2003), but offers vertical averages of Earth's atmosphere for global-, hemispheric-, and gridded time series, with no geographical gaps aside from the extreme polar regions.

Temperature time series from different data sources can be used to test the robustness of conclusions about historical climate change. Each dataset has weaknesses and strengths, and is subject to corrections and adjustments that can alter details of climate history (Brohan and others, 2006; Thompson and others, 2008) and resolve major disagreements over climate trends (Sherwood and others, 2005; Christy and others, 2007; Santer and others, 2008). Some differences between data sets are specious, such as the warmest-year-in-history honor between the HadCrut3 (warmest=1998) and GISTEMP (warmest=2005) data sets, but others reveal substantive interactions within Earth's climate system, as we argue below.

Finally, atmosphere-ocean  $CO_2$ -exchange is facilitated both by SST and the thickness of the mixed layer (Takahashi and others, 2002). The North Atlantic Ocean is a significant carbon sink whose mixed-layer thickness increases with surface-wind intensity. Storminess in the North Atlantic and other high-latitude regions may correlate better than local SST with  $CO_2$  fluctuations, and offer further insight into carbon-cycle dynamics. We use a time series of the North Atlantic Oscillation (Hurrell and Deser, 2009) since 1950 available from the NOAA Climate Prediction Center (http://www.cpc.noaa.gov/). We also obtained a time series since 1951 of the Southern Oscillation Index from the same NOAA website. Both NAO and SOI time series are based on normalized pressure differences between fixed points on the globe, and are proxies for large-scale atmospheric circulation patterns.

### METHODS

We employ multiple-taper (MT) spectrum estimation (Thomson, 1982; Prieto and others, 2009), which suits coherence estimation well. Multitaper algorithms compute a finite set of statistically independent spectrum estimates  $X^{(k)}(f)$ , called eigenspectra, from a narrow spectral band about a target frequency f. The degrees-of-freedom (dof) in a coherence estimate C(f) determine its confidence levels for nonrandomness (Abramowitz and Stegun, 1965; Thomson and others, 2007). The dof are determined by the time-bandwidth product p, which scales the estimation half-bandwidth  $\Delta f = pf_{R}$ , where  $f_R = 1/T$  is the Rayleigh frequency of a time series with duration T. For a 50-year series,  $f_R = 0.02$  cpy. For time-bandwidth product p, there are 2p - 1 orthogonal Slepian data tapers  $w^{(k)} = \{w_0^{(k)}, w_1^{(k)}, \ldots, w_{N-1}^{(k)}\}$  that have useful spectral leakage properties. Given two data series  $x = \{x_0, x_1, \ldots, x_{N-1}\}$  and  $y = \{y_0, y_1, \ldots, y_{N-1}\}$  we compute spectrum estimates

$$X^{(k)}(f) = \sum_{n=0}^{N-1} w_n^{(k)} x_n \exp(i2\pi f n \Delta t)$$
  
$$Y^{(k)}(f) = \sum_{n=0}^{N-1} w_n^{(k)} y_n \exp(i2\pi f n \Delta t)$$
(1)

for Slepian tapers indexed by k = 0, 1, ..., K - 1, where  $K \le 2p - 1$ . The coherence C(f) is estimated via

$$C(f) = \frac{\sum_{k=0}^{K-1} (X^{(k)}(f)) * Y^{(k)}(f)}{\sqrt{\sum_{k=0}^{K-1} |X^{(k)}(f)|^2} \sqrt{\sum_{k=0}^{K-1} |Y^{(k)}(f)|^2}}$$
(2)

The mod-squared coherence  $|C(f)|^2$  will have statistics related to the *F* varianceratio test with 2 and (2K - 2) degrees of freedom. The correlation transfer function (Park and Levin, 2000)

$$H(f) = \left(\sum_{k=0}^{K-1} (X^{(k)}(f)) * Y^{(k)}(f)\right) / \left(\sum_{k=0}^{K-1} |X^{(k)}(f)|^2\right)$$
(3)

assumes that the process  $\{x_n\}$  is a causal factor in the process  $\{y_n\}$ . Switching the variables X and Y reverses the assumed direction of causality. The online supplement to Park (2009) provides test data and computer code for some simple coherence and transfer function relationships: time lags, time leads, integral and derivative relationships, *et cetra*.

Both temperature and  $CO_2$  time series have significant autoregressive memory and red spectra. We prewhiten all time series with a simple one-month first-difference (fig. 1). We difference the quasi-weekly GLOBALVIEW time series after shifting by four samples instead of one to preserve spectral relationships. For each time interval considered, we subtract the mean of the first-differenced series, which removes any original trend. We fit for periodic components at f = 1 and f = 2 cpy and subtract them. Because the yearly cycles of climate observations have some variability, the amplitudes of the fitted f = 1 and f = 2 cpy cosinusoids depend on the time interval studied. For 30-yr data series, we choose p = 3.6 and K = 6 Slepian tapers for spectrum estimation, so that the estimation half-bandwidth  $\Delta f = 0.12$  cpy.

We computed a wavelet-based coherence estimator for selected data-series pairs, in order to investigate the transition in the coherence relationship reported by Park (2009). Similar to multiple-taper Fourier techniques, one can define a frequency-



Fig. 1. Data used in this study. Upper: Global-average temperature-anomaly data from the Hadley Centre database (Brohan and others, 2006) since 1850 and  $CO_2$  measured at Mauna Loa, Hawaii since 1958 (Keeling and others, 2009). Data are sampled monthly. Lower: Same datasets 1958–2008, after first-difference filter, demeaning, and subtraction of annual and semiannual cycles.

concentrated sequence of Slepian wavelets (Lilly and Park, 1995) that can be used for spectrum estimation. Multi-wavelet coherence has been applied to both seismology (Bear and Pavlis, 1999) and climate (Park and Mann, 2000; Koch and others, 2003) data analysis. The Slepian wavelets optimize spectral-leakage resistance in a user-defined bandwidth about a carrier frequency  $f_o$ , within a time interval that corresponds to a fixed number of cycles of period  $T_o = 1/f_o$ . The advantage of wavelets over so-called "evolutive" spectrum estimates is that coherence can be estimated over a fixed number of cycles, rather than a fixed time interval. Aside from this distinction, the interpretation of wavelet-based coherence differs little from standard Fourier-transform approaches. In this study, we utilize Slepian wavelets of 4-cycle duration, time-bandwidth product p = 2.5, and use K = 3 wavelets to estimate time-varying spectra and coherence.

We also apply multivariate coherence estimators, in addition to coherence estimates between pairs of data series. Because the Slepian tapers are mutually orthogonal, the complex-valued eigenspectral estimates  $Y^{(k)}(f)$  are statistically independent for white noise time series, a property that can be exploited even where the underlying spectra are smoothly-varying, that is, locally white, in the frequency domain (Thomson, 1982). To discover band-limited correlated variability in a multivariate dataset of M data series, where M > K, we assemble an  $M \times K$  matrix **A**, where  $A_{mk} = Y_m^{(k)}(f)$  is the *k*th eigenspectrum estimate at *f* of the *m*th data series. We compute the singular-value decomposition

$$\mathbf{A} = \sum_{k=1}^{K} \lambda_k \mathbf{u}_k \otimes \mathbf{v}_k^* \tag{4}$$

where the singular values  $\lambda_1 \geq \lambda_2 \cdots \lambda_K \geq 0$  and the asterisk denotes complex conjugation (Mann and Park, 1999; Apipattanavis and others, 2009). The left-singular vector  $\mathbf{u}_k$  represents the relative amplitudes and phases of the *M* data series in the *k*th mode of correlated variability, that is, the *k*th empirical orthogonal function (EOF). The right-singular vector  $\mathbf{v}_k$  represents the narrowband spectral signature of the *k*th EOF.

Even the most sophisticated statistical methods can fail if their starting assumptions are incorrect. Statements such as "99 percent confidence for nonrandomness" are precisely valid only in their proper context. Multiple-taper spectrum estimates assume the time series to be stochastic and stationary (Thomson, 1982), and many tests for deterministic or nonstationary signals have been devised (Thomson, 1982, 1990; Thomson and others, 2007). Thomson (1997) reported that the global-average temperature time series fails tests for stationary behavior near f = 0.25 and f = 0.5 cyc/yr, probably due to long-term trends in, respectively, ENSO and the Quasi Biennial Oscillation (QBO, see Labitzke, 2005). Wavelet methods can be used to explore such nonstationarity, as we argue below. Unexpected biases to coherence estimates are possible if the data possess strong periodic or near-periodic signals, such as the 11-year and 22-year sunspot cycles, or if calendar effects alias 7-day work-week cycle into the monthly-averaged CO<sub>2</sub> data (Cleveland and Devlin, 1980). Because statistical assumptions can rarely be guaranteed 100 percent, conclusions from this study should motivate, rather than terminate, further investigation.

# COHERENCE BETWEEN ATMOSPHERIC CO $_{\rm 2}$ AND GLOBAL TEMPERATURE

Park (2009) showed that C(f) between Hadley-Centre global-average temperature and  $CO_2$ -concentration at Mauna Loa USA over 1958 to 1988 agrees largely with Kuo and others, 1990, but coherence over 1979 to 2008 differs significantly (fig. 2). For 1958 to 1988  $|C(f)|^2$  exceeds the 99 percent confidence level for nonrandomness over 0.2 f < 0.6 cpy, breaches the 90 percent confidence level for non-periodic variability near f = 1 cpy, and breaches the 90 percent confidence level within 1.25 f < 1.6 cpy. For 1979 to 2008  $|C(f)|^2$  exceeds 99 percent confidence over a narrower low-frequency bandpass and displays near-zero coherence at f = 1.0 cpy.

The coherence phase  $\arg(C(f))$  indicates a shift in the nature of the temperature-CO<sub>2</sub> linkage as well as in its strength. During 1958 to1988 a linear phase ramp mimics the effect of a 6-month delay in CO<sub>2</sub> fluctuations relative to those in global  $\Delta T$ . Kuo and others (1990) reported a 5-month delay; the discrepancy might relate to updates in the Hadley Centre SST dataset (Brohan and others, 2006). During 1979 to 2008 the interannual  $\arg(C(f)) \sim 90^\circ$  where  $|C(f)|^2$  is significant, that is, a constant phase rather than a phase ramp. The correlation transfer function H(f) has amplitudes of 2 to 3 ppm/°C over the bandwidth of high coherence at interannual frequencies, but the apparent transfer function for annualcycle fluctuation is substantially larger: 4 to 5 ppm/°C. The 90° phase relationship and the decreasing trend in H(f) amplitude for f < 1 cpy are consistent with atmospheric CO<sub>2</sub> concentration behaving as a scaled integral of interannual global temperature fluctuations (Park, 2009).



Fig. 2. Coherence between global-average temperature (Hadley-Centre time series) and monthly Mauna Loa  $CO_2$  observations. Left panels compare squared coherence  $|C(f)|^2$  and correlation transfer function H(f) for 1958–1988 (red) and 1979–2008 (blue) time intervals. The 90% and 99% confidence intervals for 2 and 10 degrees of freedom (2 and 8 dof near the annual cycle) are plotted in green. Right panels plot the coherence phases for the two data intervals, with the phases of simple time-series relationships plotted in green. Uncertainty estimates for transfer function and coherence phase are plotted in grayshade, with lighter gray for larger uncertainties.

We supplement the Park (2009) coherence results to include global-average temperature estimates from sources beyond the Climate Research Unit at the University of East Anglia. The GISTEMP dataset of the NASA Goddard Institute for Space Studies offers a global-average data product computed from land-based surface-air temperature observations (Hansen and others, 1999), without information from SST measurements, either from ships or from satellites. The GISTEMP global-average data series coheres with the Mauna Loa CO<sub>2</sub> time series in a manner similar to HadCrut3 (fig. 3). Significant coherence at interannual frequencies is observed for both the 1958 to 1988 and 1979 to 2008 time intervals, with a marginally significant coherence (>90% confidence) in 1958 to 1988 for annual-cycle fluctuations. The correlation transfer function |H(f)| falls in the ranges defined by the HadCrut3 data set, with a smoother decline in amplitude with increasing f in the interannual bandpass. The coherence phase  $\arg(C(f))$  in figure 3 follows roughly the behavior seen in figure 2, but differences are significant, relative to the formal uncertainties on arg C(f). The expected 6-month phase ramp during 1958-1988 is choppy, and the interannual phase shift of CO<sub>2</sub> relative to GISTEMP during 1979 to 2008 is 60 to 75° rather than 90°. The latter discrepancy implies that interannual CO<sub>2</sub> fluctuations occur slightly ahead of a first-integral of GISTEMP temperature fluctuations. This can be explained if the principal source of CO<sub>2</sub>-temperature coherence is related to SST fluctuations and their effect on atmosphere-ocean gas exchange, and that the land-based GISTEMP global average data series responds to SST variations with time lags of 2 to 4 months. Coherence estimates between the HadCrut3 and GISTEMP time series (not shown)



Fig. 3. Coherence between global-average temperature (NASA-GISS time series from meteorological stations) and monthly Mauna Loa  $CO_2$  observations. Left panels compare squared coherence  $|C(f)|^2$  and correlation transfer function H(f) for 1958–1988 (red) and 1979–2008 (blue) time intervals. The 90% and 99% confidence intervals for 2 and 10 degrees of freedom (2 and 8 dof near the annual cycle) are plotted in green. Right panels plot the coherence phases for the two data intervals, with the phases of simple time-series relationships plotted in green. Uncertainty estimates for transfer function and coherence phase are plotted in grayshade, with lighter gray for larger uncertainties.

agree with this scenario. They reveal high, but not perfect coherence, that is,  $|C(f)|^2$  above the 99 percent confidence level for nonrandomness, and an irregular coherence phase  $\arg(C(f)) \sim 15$  to 30°.

Yet another estimation philosophy is represented by the MSS satellite data for lower-troposphere temperatures, in which pointwise surface temperature measurements are replaced by vertical integrals of inferred temperature in the lowest few kilometers of Earth's atmosphere (Spencer and Christy, 1990; Mears and others, 2003). Both sources of monthly global-average temperature data lead to coherence estimates with 1979 to 2008 Mauna Loa  $CO_2$  that comport with results from the HadCrut3 data set, with deviations similar to those of the GISTEMP coherence estimates (fig. 4). Unlike the surface temperatures, lower-tropospheric temperatures are marginally coherent with Mauna Loa  $CO_2$  at 0.6 f < 0.9 cpy, far from the ENSO band, with correlation transfer function amplitude near 1 ppm/°C. The tropospheric transfer function has 70 to 80° phase shift and lower amplitude in the ENSO band, compared to HadCrut3 surface measurements, consistent with a delay and a dampening of SST fluctuations with height above Earth's surface.

Differing  $CO_2$  coherence among the different temperature data series could be due to many factors, including influences that involve other climate variables, for example, precipitation. Coherence with individual gridpoint time series in the Had-Crut3 data set (see Appendix A) reveal strong coherence of  $CO_2$  with interannual temperature variations in the open tropical ocean. p $CO_2$  surveys of the world ocean (Takahashi and others, 2002, 2009) show that low-latitude p $CO_2$  in the oceans exceeds

494



Fig. 4. Coherence between global-average lower troposphere temperature (Microwave Sounding Unit satellite) and monthly Mauna Loa  $CO_2$  observations. Left panels compare squared coherence  $|C(f)|^2$  and correlation transfer function H(f) for 1979–2008 data interval obtained from Univ. Alabama-Huntsville (UAH, red) and Remote Sensing Systems (RSS, blue). The 90% and 99% confidence intervals for 2 and 10 degrees of freedom (2 and 8 dof near the annual cycle) are plotted in green. Right panels plot the coherence phases for the two data sources, with the phases of simple time-series relationships plotted in green. Uncertainty estimates for transfer function and coherence phase are plotted in grayshade, with lighter gray for larger uncertainties.

atmospheric  $CO_2$  concentrations. The tropical oceans are therefore a source of  $CO_2$  to the atmosphere; the net-positive ocean uptake of  $CO_2$  is achieved via a larger carbon sink at high latitudes. Tropical SST increments should increment the p $CO_2$  imbalance and boost  $CO_2$  flux from ocean to atmosphere in low latitudes, and figures 2 to 4 are consistent with this process. Physical controls on gas exchange are not the only factor to consider. Optical estimates of phytoplankton abundance suggest a negative correlation between tropical SST and biological utilization of dissolved  $CO_2$  (Boyce and others, 2010). This poorly-calibrated but reinforcing biologic effect would boost surface-ocean p $CO_2$  and increase  $CO_2$  flux to the atmosphere, reinforcing the effect of a temperature increase. Because the world ocean is currently a net carbon sink, a tropical SST increase effectively slows the net rate of oceanic  $CO_2$ -uptake.

### WHEN AND HOW DID THE SHIFT IN COHERENCE PHASE OCCUR?

Restricting attention to Mauna Loa  $CO_2$  and HadCrut3 global-average temperature, we apply multi-wavelet coherence analysis to investigate the evolution of interannual  $\arg(C(f))$  (Lilly and Park, 1995; Park and Mann, 2000). Data series are firstdifferenced, demeaned and decycled at f = 1 and 2 cpy before computing the Slepian wavelet transform. Figure 5 demonstrates that coherence in the time-period domain is



Fig. 5. Wavelet spectra and coherence for the Hadley-Centre global-average temperature and Mauna Loa  $CO_2$  time series. Lower panels graph wavelet spectra in grayshade. The y-axes scale with the logarithm of wavelet period, in powers of two; darker shades imply larger spectrum values. Middle panel plots wavelet coherence, with lighter shades for larger  $|C|^2$  values. For  $|C(f)|^2 > 0.7$  we plot wavelet correlation as scaled arrows. A rightward arrow implies zero phase, and an up-arrow implies a 90° phase angle. Top panel graphs the time delay implied by the correlation phase for wavelets centered on 4-yr period, near the ENSO climate resonance.

most extensive during the ENSO episodes of the 1970s and the late 1990s, near maxima in the wavelet spectra for both  $CO_2$  and temperature data series. Significant coherence is persistent for much of the 1958 to 2008 interval for wavelets at period T = 4 years, roughly the ENSO period. The phase of the wavelet coherence at 4-yr time scale, indicated by arrows at time-period points where  $|C(f)|^2 > 0.70$ , shows a noticeable shift near the year 1980. Translating this shift into a time-lag at 4-yr period (fig. 5A) the wavelet coherence phase-lag is observed to jump from ~6 months to ~12 to 15 months across a transition interval of 5 years. We observe depressed  $CO_2$ -temperature wavelet coherence during the transition. A peak 15-month phase lag for a 4-cycle (16-yr) time interval is found to be centered on the mid-1980s, after which the  $CO_2$ -temperature wavelet phase decreased to 10 to 12 months for time intervals centered in the 1990s and 2000s.

The  $\arg(C(f))$  phase would increment monotonically if it were responding to a growing saturation of the oceanic carbon sink. Rather, a marked transition in the inter-annual  $CO_2$ -temperature phase is suggested in the late 1970s. This transition coincides with the oft-cited 1970s climate transition that heralded a 20-year secular rise in temperatures (Trenberth, 1990; Graham, 1994; Park and Mann, 2000; Meehl and others, 2009), following a 35-yr interval of stagnant temperatures in the Northern Hemisphere. Because the 1970s climate transition is commonly associated with a change in ocean circulation, as well as a polarity change in the Pacific Decadal Oscillation (PDO, see Mantua and Hare, 2002) it seems likely that the CO<sub>2</sub>temperature coherence transition identified by Park (2009) is associated with this event. Park and Mann (2000) found, via a multiwavelet analysis of an earlier version of the HadCrut gridded temperature data product (Jones and Briffa, 1992) that the spatio-temporal patterns of multicycle 20th-century ENSO episodes were multifarious, displaying global-average temperature fluctuations that varied greatly in amplitude from decade to decade. It is therefore not surprising that CO<sub>2</sub>-temperature coherence phase shifted abruptly between the two largest ENSO episodes of the late 20th century.

### COHERENCE BETWEEN ATMOSPHERIC CO<sub>2</sub>, NAO AND SOI

Oceanic  $CO_2$  uptake is large in the North Atlantic (Takahashi and others, 2009), but we do not find pervasive interannual coherence between atmospheric  $CO_2$  time series and gridpoint temperatures in the region. Thickening of the surface mixed layer via wind stress is hypothesized to be more important than SST for  $CO_2$ -uptake in the North Atlantic, and the North Atlantic Oscillation, NAO, a normalized sea-level pressure index, correlates with storminess in the region (Hurrell and Deser, 2009). Because ENSO couples interannual temperature and circulation fluctuations, one expects to find coherence between the pressure-based Southern Oscillation Index (SOI) and atmospheric  $CO_2$  with significance comparable to C(f) with temperature time series. Unique features in SOI- $CO_2$  coherence amplitude and phase may offer clues to causal relations.

Figures 6A and 6B show the coherence between NAO and SOI is weak at most interannual frequencies during 1979 to 2008, demonstrating their status as independent climate proxies. To test overall coherence of NAO and SOI with CO<sub>2</sub> measurements, we averaged  $|C(f)|^2$  for all 14 near-continuous GLOBALVIEW stations (fig. 6C). The average  $|C(f)|^2 \sim 0.5$  for SOI for 0.2 f < 0.5 cpy. Because the significance thresholds for multiple CO<sub>2</sub> records are lower than for single records (28 and 140 dof versus 2 and 10 dof), the average  $|C(f)|^2$  for SOI exceeds the 99 percent confidence limit for nonrandomness for 0.15 f < 0.8 cpy, a broader bandpass than typical for ENSO variability.

By contrast, average  $|C(f)|^2$  for NAO fails to exceed 90 percent confidence for nonrandomness over most f < 1.0 cpy. The largest peak in average  $|C(f)|^2$  for NAO occurs near f = 0.35 cpy, but near this frequency the NAO is strongly coherent with the



Fig. 6. Coherence between monthly values of the Southern Oscillation Index (SOI) and the North Atlantic Oscillation (NAO) and 14 long-running CO<sub>2</sub> GLOBALVIEW sites. (A) Interannual  $|C(f)|^2$  between NAO and SOI for 1979–2008. The 90% and 99% confidence intervals for 2 and 10 degrees of freedom (2 and 8 dof near the annual cycle) are plotted in green. (B) The coherence phase  $\arg(C(f))$  between NAO and SOI, with grayshade uncertainties. (C) Summed squared coherences  $|C(f)|^2$  between 14 GLOBALVIEW data series versus SOI (red) and NAO (blue). The 90% and 99% confidence intervals for 28 and 140 degrees of freedom (28 and 112 dof near the annual cycle) are plotted in green. (D) Summed squared coherences  $|C(f)|^2$  between NAO and 14 GLOBALVIEW data series after subtraction of the CO<sub>2</sub> variance coherent with SOI. The 90% and 99% confidence intervals for 28 and 84 dof near the annual cycle) are plotted in green.

SOI. We subtracted the portion of the GLOBALVIEW  $CO_2$  spectra that are coherent with the SOI and recomputed the coherence of the residuals with the NAO (fig. 6D). The coherence peak near f = 0.35 cpy disappears for this case, leaving only two  $|C(f)|^2$ peaks that broach the 99 percent confidence level, near the annual frequency f = 1 cpy and at f = 0.13 cpy. Mann and Park (1994, 1996) found a mode of correlated variability near f = 0.13 cpy in global-gridded temperatures and in joint temperature-pressure data sets at 95 percent confidence for nonrandomness. Mann and Park (1994) reported that the spatial pattern of this 7 to 8-yr period mode of variability resembled the quasi-biennial oscillation, QBO, centered around the North Atlantic. This corroboration suggests that the f = 0.13 cpy peak in average  $|C(f)|^2$  between NAO and the GLOBALVIEW  $CO_2$  data for 1979 to 2008 represents a meaningful natural climate process.

The correlation amplitudes and phases at selected frequencies [see eq (3)] lend support to meaningful correlation between SOI, NAO and CO<sub>2</sub> fluctuations (fig. 7). Phased coherence C(f) for SOI at the ENSO frequency f = 0.25 cpy shows a consistent pattern, except for outlier station **cmn** (fig. 7B). At f = 0.13 cpy, NAO fails to cohere with two stations, **brw** (Barrow, Alaska) and **nwr** (Niwot Ridge, Colorado), but exhibits consistent phase angles at the remaining stations (fig. 7C). Near-identical correlation coefficients at interannual frequencies (figs. 7B and 7C) seem consistent with a widespread interannual causal factor, that is, the ocean. Long-term CO<sub>2</sub> fluctuations



Fig. 7. Geographic correlation pattern between NAO, SOI and 14 GLOBALVIEW CO<sub>2</sub> sites for 1979–2008 at selected frequencies: (A) locations of 14 GLOBALVIEW stations; (B) transfer function for SOI at f = 0.25 cpy. Length of arrow scales correlation value. Darker arrow shades indicate higher coherence. Circles indicate uncertainty of transfer function. Eastward arrow is 0° phase. Upward arrow is 90° phase. (C) Transfer function for NAO at f = 0.13 cpy. (D) Transfer function for SOI at f = 1.00 cpy. (E) Transfer function for NAO at f = 1.00 cpy.

are hypothetically linked to North Atlantic Deep Water formation as part of the conveyorbelt hypothesis for Pleistocene glacial cycles (Maasch and Saltzman, 1990). Our result suggests that fluctuations in this process may be observable on decadal time scales.

At f = 1 cpy, both SOI and NAO appear to cohere significantly with the residual CO<sub>2</sub> variability after a fixed annual cycle is subtracted, with the largest correlation coefficients at GLOBALVIEW stations in North America and Europe (figs. 7D and 7E). Correlation phases for SOI for annual-cycle variability are near-parallel. Correlation phases for NAO alternate polarity across continents and/or oceans. The diverse sizes of correlation coefficients for the f = 1 cpy residuals argue for a more localized annual-cycle causal factor, for example, the Northern-Hemisphere forests.

The coherence of SOI with both global temperatures and with the Mauna Loa  $CO_2$  record exhibits differing phase lags between the mid- and late-20th century, consistent with a major climate shift in the late 1970s. The SOI tends to be negatively correlated with tropical SST fluctuations (McPhaden and others, 2006) and during 1979 to 2008 it also exhibited a phase advance of roughly 6 months for 0.2 f < 0.5 cpy (fig. 8). The phase ramp deviates for f > 0.5 cpy, outside the ENSO bandpass. Phase complexities are absent for coherence between HadCrut3 temperatures and SOI for the 1951 to 1980 interval (fig. 9), for which the interannual  $\arg(C(f)) \sim 150^{\circ}$ . We



# HadCrut3 global temp vs. SOI (1979-2008)

Fig. 8. Coherence between monthly global-average temperature (Hadley-Centre values) and the Southern Oscillation Index (SOI) for 1979–2008. Upper: interannual  $|C(f)|^2$ , with the 90% and 99% confidence intervals for 2 and 10 degrees of freedom (2 and 8 dof near the annual cycle) plotted in green. Center: Transfer-function amplitude with grayshade uncertainties. Lower: Coherence phase with grayshade uncertainties. SOI is negatively correlated with global-average temperature, but deviates from a 180° phase lag along a phase ramp (green line) that indicates a 6-month time lead for SOI.



# HadCrut3 global temp vs. SOI (1951-1980)

Fig. 9. Coherence between monthly global-average temperature (Hadley-Centre values) and the Southern Oscillation Index (SOI) for 1951–1980, mainly prior to the late 1970s climate transition. Upper: interannual  $|C(f)|^2$ , with the 90% and 99% confidence intervals for 2 and 10 degrees of freedom (2 and 8 dof near the annual cycle) plotted in green. Center: Transfer-function amplitude with grayshade uncertainties. Lower: Coherence phase with grayshade uncertainties. SOI is approximately negatively correlated with global-average temperature, but follows a 150° phase lag (green line) rather than 180°.

conclude that the late-1970s climate transition altered the correlation phase of SOI and ENSO-related SST fluctuations.

The coherence of SOI with Mauna Loa  $CO_2$  also undergoes a shift between the 1958 to 1988 and 1979 to 2008 time intervals (figs. 10 and 11). In the ENSO bandpass (0.2-0.5 cpy) the shift does not have a simple time-domain analog. However, outside the ENSO bandpass the SOI coherence with Mauna Loa  $CO_2$  has simpler phase, in particular,  $\arg(C(f)) \sim 0^\circ$  for 0.5 f < 0.75 cpy, even where 90 to 99 percent confidence for nonrandomness is not satisfied. We conclude that Mauna Loa  $CO_2$  correlates more simply with temperature in the interannual ENSO bandpass, and more simply with pressure-based climate variables both above (via SOI) and below (via NAO) the ENSO bandpass.

Coherence with the annual-cycle residual switches from  $\arg(C(f)) \sim -90^{\circ}$  in 1958 to 1988 to  $\arg(C(f)) \sim 0^{\circ}$  in 1979 to 2008. This last transition suggests a positive correlation between SOI and CO<sub>2</sub> annual-cycle fluctuations before the late 1970s climate transition, and a negative correlation between SOI and the time-derivative of the CO<sub>2</sub> annual-cycle residual after the transition. Even though the quasi-annual correlation amplitude is relatively large and the phase relation simple, a causal factor is not obvious to us.

### DISCUSSION

Occam's Razor is not a perfect tool, but a focus on the simplest observational relationships is often useful for generating and testing hypotheses. We have used the tools of spectrum analysis to explore relationships between Earth's carbon cycle and its climate system on time scales of one year and longer. The coherence C(f) tests whether two time series are correlated near frequency f to varying degrees of statistical confidence. The coherence phase  $\arg(C(f))$  can test causal relationships between time series. Simple phase ramps and phase shifts in the frequency domain correspond to simple models in the time domain that are not always detected with time-domain correlation techniques. Computer codes that demonstrate these relationships can be downloaded from the online supplement to Park (2009). For instance, a phase ramp in the frequency domain corresponds to a lead or lag in the time domain, and this behavior can be detected with a lagged correlation in the time domain. However, a  $90^{\circ}$ phase relationship will likely be undetected in the time domain, because a broadband signal and its derivative often do not have the same shape. In addition, a coherence relationship that changes with frequency is more difficult to detect with a time-domain correlation, which typically is dominated by the time scale with largest amplitude.

Our analysis demonstrates subtle, but significant, changes in the relationship between Earth's climate and atmospheric CO<sub>2</sub> in the late 20th century, extending the results of Park (2009). The earlier study found a shift in the coherence between the monthly global-average Hadley Centre temperature-anomaly data set (Brohan and others, 2006) and monthly Mauna Loa CO<sub>2</sub> observations (Keeling, 2008). Park (2009) hypothesized that a transition from a 6-month time lag (for 1958-1988) to a 15-month time lag (for 1979-2008) reflected a more-sluggish response of atmosphere-ocean CO<sub>2</sub>-exchange to SST fluctuations in the late 20th century, supporting model-based estimates of a weakening oceanic carbon sink (Le Quere and others, 2007; Schuster and others, 2009; Khatiwala and others, 2009); see Knorr (2009) for a dissenting view. In this study we corroborate the transition, but use wavelet coherence to reject a gradual transition via oceanic saturation. We infer instead a rapid transition coinciding with the late-1970s climate shift near the most-recent inflection point for secular change in the global-average temperature (Trenberth, 1990; Graham, 1994; Park and Mann, 2000; Meehl and others, 2009). Parallel coherence estimates using the pressurebased Southern Oscillation Index (SOI) corroborate the notion of a carbon-cycle shift coeval with the climate shift, and add detail to the picture.



# SOI vs. Mauna Loa CO<sub>2</sub> (1979-2008)

Fig. 10. Coherence between the Southern Oscillation Index (SOI) and monthly Mauna Loa CO<sub>2</sub> observations for 1979–2008. Upper: interannual  $|C(f)|^2$ , with the 90% and 99% confidence intervals for 2 and 10 degrees of freedom (2 and 8 dof near the annual cycle) plotted in green. Center: Transfer-function amplitude with grayshade uncertainties. Lower: Coherence phase with grayshade uncertainties. SOI is positively correlated with Mauna Loa CO<sub>2</sub> at both interannual and annual frequencies, but deviates from a 0° phase lag in the ENSO bandpass.



SOI vs. Mauna Loa CO<sub>2</sub> (1958-1988)

Fig. 11. Coherence between the Southern Oscillation Index (SOI) and monthly Mauna Loa  $CO_2$  observations for 1958–1988. Upper: interannual  $|C(f)|^2$ , with the 90% and 99% confidence intervals for 2 and 10 degrees of freedom (2 and 8 dof near the annual cycle) plotted in green. Center: Transfer-function amplitude with grayshade uncertainties. Lower: Coherence phase with grayshade uncertainties. SOI is positively correlated with Mauna Loa  $CO_2$  at interannual frequencies, but deviates from a 0° phase lag in the ENSO bandpass. At 90% confidence for nonrandomness, SOI exhibits a  $-90^\circ$  phase shift for coherence of annual-cycle fluctuations, with large-amplitude transfer function.

Jeffrey Park

In contrast to many studies in which interannual  $CO_2$  fluctuations are explained primarily by means of a fluctuating appetite for carbon in the terrestrial biosphere (Lee and others, 1998; Bousquet and others, 2000; Zeng and others, 2005; Buermann and others, 2007; Qian and others, 2008), we argue that oceanic climate and carboncycle exchanges are a principal driver. By "a principal driver" we do not mean the ocean is the only agent, because terrestrial  $CO_2$  flux is clearly evident in the spatial variation of data coherences described in Appendix A, and modeling studies predict substantial fluctuations in the terrestrial carbon cycle. However, temperature data restricted to continental interiors correlate more weakly with interannual  $CO_2$  variability, as illustrated in Appendix A.

Evidence for simple coherence between interannual CO<sub>2</sub> variations and ocean climate indices includes (1) a simple ramp in the coherence phase  $\arg(C(f))$  with Hadley-Centre global-average temperatures during 1958 to 1988 at frequencies  $f \approx 0.5$ cpy, consistent with a 6-month lagged response; (2) a  $90^{\circ}$ -phase in the 1979 to 2008 interval, consistent with a direct correlation between oceanic SST and CO<sub>2</sub> rate-ofchange; (3) the transition between these two relationships occurs near a recognized shift in Pacific ocean circulation in the late 1970s; (4) significant coherence between interannual CO<sub>2</sub> fluctuations at many GLOBALVIEW stations and low-latitude oceanic gridpoint temperatures in the Hadley-Centre dataset, with consistent correlation phases among different  $CO_2$  stations; (5) coherences at >90 percent confidence for nonrandomness between the pressure-related SOI and NAO indices in frequency bands above and below, respectively, the ENSO bandpass; (6) zero-phase coherence for SOI above the ENSO band, and a geographic consistency in NAO coherence phase at 7-yr period, which is a recognized decadal mode of temperature variability; (7) when correlations are sought between  $CO_2$  and global-temperature time series that do not involve sea-surface temperatures explicitly, based either on weather station (NASA-GISS) or satellite lower-troposphere (MSU) measurements, the coherence phase is perturbed, consistent with a downstream response to SST.

We suggest that oceanic control of  $CO_2$  fluctuations relates to SST and wind-stress modulation of the mixed-layer thickness in different frequency passbands. In the ENSO band the simpler coherence-phase with  $CO_2$  occurs for global-average temperature, not the pressure indices, pointing toward SST as the primary agent of  $CO_2$  flux tied to ENSO, rather than pressure-induced atmospheric circulation changes.

We find scattered and inconsistent evidence for the coherence of interannual CO<sub>2</sub> fluctuations with continental temperature, in the Hadley-Center gridpoint time series. The terrestrial carbon cycle may be influenced more by drought than by temperature, however. Interannual drought fluctuations are closely tied to ENSO, and therefore closely tied to SST (Apipattanavis and others, 2009). The principal component of CO<sub>2</sub> variability at f = 0.25 cpy, in the 1979 to 2008 GLOBALVIEW marine-boundary-layer data product (see Appendix A) has peak amplitude in the mid- to high-latitude Northern Hemisphere with a 4-month time lag relative to the tropics and the Southern Hemisphere. Park (2009) noted a similar phase relationship that contrasted interannual  $CO_2$  fluctuations at low latitude Mauna Loa, Hawaii from those at high-latitude Barrow, Alaska. The amplitude and phase lag argue for a prominent contribution by the extratropical terrestrial biosphere, but as a lagged response to oceanic SST variations. Our analysis does not exclude the possibility that the tropical terrestrial biosphere plays a major role in interannual  $CO_2$  fluctuations via drought and fire (Page and others, 2002; McPhaden and others, 2006) but terrestrial influence would need to combine with oceanic low-latitude carbon-cycle processes to produce simple coherence relationships with oceanic climate indices.

Knowing where carbon moves in the natural system is a pre-requisite to forecasting mankind's inevitable meddling in the cycle. Statistical studies like this cannot uniquely determine the causes of the correlations that are detected. Rather, they serve as experimental data for the falsification of scientific hypotheses, in particular the predictions of global carbon-cycle models. Because the natural variability of climate and of climate models is substantial, long-term statistical relationships are better tools for validating theoretical relationships than are curve-fitting experiments that involve a few years of data. Now that  $CO_2$  time series extend past a few decades, such statistical relationships can be better constrained.

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#### Appendix A

### SPATIAL PATTERNS OF $\mathrm{CO}_2\text{-}\mathrm{TEMPERATURE}$ COHERENCE IN THE ENSO BAND

Park (2009) showed that interannual coherence between  $CO_2$  and global-average temperature can be found at widely separated locations, Mauna Loa, Hawaii; Barrow, Alaska; South Pole, but with differences in transfer-function amplitude and phase. For coherence since 1979, the GLOBALVIEW data product (Masarie and Tans, 1995; GLOBALVIEW-CO<sub>2</sub>, 2009) offers a wider selection of station locations to explore these statistical relationships. We restrict attention to 14 recording locations where  $CO_2$  data series have data gaps, interpolated via the GLOBALVIEW averaging procedure, that total no more than two years out of the 30 years possible (table 1). Of these data series, only one fails to exhibit significant coherence with the HadCrut3 global-average temperature series at interannual frequencies f < 0.5 cpy: station **cmn**, a high-altitude site at Mt. Cimone, Italy.

We show station triplets in common latitude ranges. Figure 12 shows low- to mid-latitude Northern-Hemisphere stations. Figure 13 shows high-latitude Northern-Hemisphere stations. Figure 14 shows Southern-Hemisphere stations. With some variability, the GLOBALVIEW stations in figures 12 and 14 show significant coherence with global temperaure, often exceeding 99 percent confidence for non-randomness, for 0.15 f < 0.5 cpy, with coherence phase close to 90°. High-latitude NH stations show similar  $|C(f)|^2$  behavior, with a modest, but significant, phase shift. At frequencies f > 0.5 cpy significant coherence is either absent or

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Station	Lat	Long	Elev (m)	Location
alt	82.45	-62.51	210	Alert, Nunavut, Canada
asc	-7.92	-14.42	54	Ascension Is., UK
cmn	44.18	10.70	2165	Mt. Cimone, Italy
brw	71.32	-156.61	11	Barrow, AK, USA
gmi	13.43	144.78	6	Guam, USA
key	25.67	-80.16	3	Key Biscayne, FL, USA
kum	19.52	-154.82	3	Cape Kumukahi, HI, USA
mlo	19.54	-155.58	3397	Mauna Loa, HI, USA
nwr	40.05	-105.58	3526	Niwot Ridge, CO, USA
psa	-64.92	-64.00	10	Palmer Station, Antarctica
smo	-14.25	-170.56	42	Tutuila, American Samoa
bhd -	-41.41	174.87	80	Baring Head Station, NZ
spo	-89.98	-24.80	2810	South Pole, Antarctica
stm	66.00	2.00	5	Ocean Station M, Norway

 TABLE 1

 Long-running CO<sub>2</sub> observation sites in the GLOBALVIEW-CO<sub>2</sub> data product



Fig. 12. Coherence between monthly global-average temperature (Hadley-Centre time series) and quasi-weekly  $CO_2$  observations at low- and mid-latitude GLOBALVIEW stations **kum** (Cape Kumakahi, Hawaii; red), **gmi**, (Guam, USA; blue) and **key** (Key Biscayne, Florida; black). The 90% and 99% confidence intervals for 2 and 10 degrees of freedom (2 and 8 dof near the annual cycle) are plotted in green. Upper-left panel compares squared coherence  $|C(f)|^2$  for the 1979–2008 time interval. Surrounding panels plot the coherence phases for the three stations, with the phases of simple time-series relationships plotted in green. Uncertainty estimates for coherence phase are plotted in grayshade, with lighter gray for larger uncertainties.

exhibits diverse values of coherence phase (stations **key**, **alt**, **brw**, **smo**, and **bhd**). This frequency dependence reinforces the notion that interannual variations in CO<sub>2</sub> are causally related to ENSO variability.

Park (2009) estimated coherence between Mauna Loa  $CO_2$  and gridpoint temperature anomaly data over the time period 1958 to 2008, with largest  $|C(f)|^2$  in the tropical oceans. Figure 15 shows similar behavior for widely-separated stations **alt** (Alert, Canada), **key** (Key Biscayne, USA) and **psa** (Palmer Station, Antarctica) for the period 1979 to 2008 at the ENSO frequency f = 0.25 cpy. The highest coherences in the tropical ocean are for **psa**, with values breaching 99 percent confidence for nonrandomness clustered broadly in the western Pacific and Indian oceans, which comprise opposing loci of large ENSO SST fluctuations. Although  $CO_2$  data from both stations **key** and **alt** exhibit significant coherence with tropicalocean gridpoints as well, it is smaller, with an altered geographical pattern. In addition, significant  $|C(f)|^2$  for **key** and **alt** is found for clusters of gridpoints in the Northern Hemisphere, such as high-latitude North America (for **key**) and the North Pacific (for **alt**). Carbon exchange in the North Pacific switches sign both laterally and during the annual cycle (Takahashi and others, 2009, see figs. 9 and 15), and therefore may be more sensitive to interannual SST fluctuations than the high-latitude North Atlantic, where persistent  $CO_2$ -uptake is influenced more by storminess fluctuations.

Figure 16 shows the phased transfer function between  $CO_2$  and the HadCrut3 gridded temperature time series at the ENSO frequency f = 0.25 cpy. The arrows point east for positive correlation, west for negative correlation, and north for a 90° phase shift in the transfer function. Note that, despite station-bystation variations in coherence  $|C(f)|^2$  among gridpoints in the HadCrut3 data set, the geographical patterns in the transfer function H(f) are similar among stations **alt**, **key** and **psa**. For example, phase shifts near 90° characterize transfer functions H(f) in the equatorial Indian Ocean, the Caribbean Sea and in the equatorial Atlantic off Brazil. Although  $|C(f)|^2$  is significant in the climatically-sensitive eastern tropical Pacific, the transfer function H(f) is small for gridpoints along the equator itself.

High  $|C(f)|^2$  between ocean SST and atmospheric CO<sub>2</sub> suggests that ocean-atmosphere CO<sub>2</sub> flux is the determining factor in interannual CO<sub>2</sub> variations, but correlated factors are also relevant, such as precipita-



Fig. 13. Coherence between monthly global-average temperature (Hadley-Centre time series) and quasi-weekly  $CO_2$  observations at high-latitude GLOBALVIEW stations **stm** (Station M, Norway; red), **alt** (Alert, Canada; blue) and **brw** (Barrow, Alaska; black), for the 1979–2008 time interval. Plotting conventions identical to figure 12.



Fig. 14. Coherence between monthly global-average temperature (Hadley-Centre time series) and quasi-weekly CO<sub>2</sub> observations at Southern-Hemisphere GLOBALVIEW stations **smo** (Samoa, USA; red), **bhd** (Baring Head, New Zealand; blue) and **psa** (Palmer Station, Antarctica; black), for the 1979–2008 time interval. Plotting conventions identical to figure 12.



Fig. 15. Coherence between monthly gridpoint temperature anomalies (Hadley-Centre time series) and quasi-weekly CO<sub>2</sub> observations at GLOBALVIEW stations **alt** (Alert, Canada), **key** (Key Biscayne, Florida) and **psa** (Palmer Station, Antarctica), at ENSO bandpass f = 0.25 cpy for the 1979–2008 time interval. Symbol size proportional to coherence  $|C(f)|^2$ , with black circles >99% confidence for nonrandomness, blue circles >95% confidence, red circles >90% confidence, green circles <90% confidence.



Fig. 16. Correlation transfer function between monthly gridpoint temperature anomalies (Hadley-Centre time series) and quasi-weekly  $CO_2$  observations at GLOBALVIEW stations alt (Alert, Canada), key (Key Biscayne, Florida) and psa (Palmer Station, Antarctica), at ENSO bandpass f = 0.25 cpy for the 1979–2008 time interval. Symbol size proportional to transfer function magnitude. |H(f)|, with rightward arrows at 0° phase, upward arrows at 90° phase. Color-coding follows that of figure 15.

# CO<sub>2</sub> ENSO Cycle (0.25 cyc/yr)



Fig. 17. Empirical orthogonal functions (EOFs) for ENSO-band  $CO_2$  variability (f = 0.25 cpy) in the marine-boundary-layer (MBL) GLOBALVIEW atmospheric- $CO_2$  data product, which comprises 41 time series, interpolated and extrapolated from fixed-station, aircraft and ship data at evenly-spaced values of sin(latitude) (GLOBALVIEW-CO<sub>2</sub>, 2009). Upper panel graphs the amplitude of each time series in the primary and secondary EOFs versus latitude (<0 in the Southern Hemisphere). The lower panel graphs the relative phasing of the MBL time series within the EOF. The primary EOF increases amplitude in steps by a factor of two from the southern to northern polar regions, with only modest change in phase. A phase increment of ~30° occurs near 40°N latitude, corresponding to a 4-month phase lag in mid- to high-latitude Northern Hemisphere. The secondary EOF contributes a small proportion of the f = 0.25 cpy cycle at all latitudes.

tion, drought and fire on land (Rodenbeck and others, 2003; McPhaden and others, 2006). Tropical drought and fire have been highlighted in associating  $CO_2$  increases with the 1997 to 1998 ENSO (Page and others, 2002). However, the largest geographical  $CO_2$ -variation in the GLOBALVIEW data set identifies the northern midlatitude continents for significant ENSO-related carbon-budget impacts. We performed a frequency-domain singular-value decomposition (SVD) on the GLOBALVIEW  $CO_2$  marine-boundary-layer (MBL) data product to investigate global correlations. Figure 17 shows the latitude dependence of the first

two principal components at f = 0.25 cpy. The first principal component has 2.5 to 5 times the amplitude of the second component at all latitudes, and has only modest variation in phase, confirming the consistent behavior seen among the near continuous data series from individual stations. In low latitudes where high coherence with gridpoint temperatures is evident (fig. 15) the peak-to-peak amplitude of the signal is approximately 0.75 ppm CO<sub>2</sub>, declining to 0.50 ppm CO<sub>2</sub> at high southern latitudes. This decline is crudely consistent with a signal that is generated in the tropics and transported poleward via atmospheric mixing. In the northern high latitudes, however, the principal component amplitude climbs toward 1.0 ppm CO<sub>2</sub>, as well as exhibiting a localized increase in phase between 30°N and 45°N.

The localized phase gradient and the amplified signal argue against simple transport of tropical air toward the northern high latitudes. Despite exhibiting significant coherence with gridpoint temperature time series in low latitudes, CO<sub>2</sub> fluctuations north of 45°N likely are responding to a correlated extratropical climate process, that is, an ENSO teleconnection. Drought-induced fires seem unlikely, because such fires are documented mainly in tropical and southern-hemisphere land areas subject to deforestation pressure. However, mild interannual drought in the mid- to high-latitude terrestrial biosphere that effects 10 to 20 percent changes in seasonal carbon uptake could explain this signal. Apipattanavis and others (2009) found joint modes of variability between oceanic SST and the Palmer Drought Severity Index at several frequencies within the ENSO band during the 20th century at the 90 to 95 percent confidence level for nonrandomness. Figure 7B of Apipattanavis and others (2009) graphs the principal singular vector  $\mathbf{u}_1$  from equation (4) above. Their spectral-domain EOF at ENSO periods has substantial amplitude at gridpoints in the mid- and high-latitude Northern Hemisphere, verifying a correlation between tropical SST fluctuations and extratropical drought severity.

#### Appendix B

### A FAILED HYPOTHESIS TEST FOR QUASI-ANNUAL COHERENCE

Are the local maxima in the  $CO_2$  wavelet spectrum near T = 1-yr target period (fig. 5) related to the marginal quasi-annual coherence between Mauna Loa CO<sub>2</sub> and global average temperature during 1958 to 1988 (figs. 2 and 3)? Close inspection suggests no. Figure 5 shows intervals of notable spectrum amplitude near 1-year period for the Mauna Loa CO<sub>2</sub> time series near the beginning and end of the record, but this behavior is less evident for the HadCrut3 wavelet spectrum. Some of the inverted-triangle features in the wavelet spectra resemble pulse-like signals, such as P- and S-wave arrivals in seismograms (Lilly and Park, 1995).

One of the few intervals of significant wavelet coherence near 1-yr period occurs proximal to the 1991 eruptions of the Pinatubo and Cerro Hudson volcanoes, two of the largest eruptions of the last 50 years (LeGrand and Wagenbach, 1999). The phase of the coherence near 1990 is consistent with global cooling following a rise in CO<sub>2</sub>, the opposite of a greenhouse response, but the cause of cooling after Pinatubo is sulfate aerosols from the eruption, not its  $CO_2$  emission (Angell, 1993; Robock, 2000). The observed  $CO_2$ fluctuations could be due to the transient impacts of aerosols on yearly plant growth via lower temperatures (Lucht and others, 2002) or scattered-light photosynthesis (Gu and others, 2003), rather than a direct effect of volcanic CO<sub>2</sub> emissions. However, the wavelet spectrum of Mauna Loa CO<sub>2</sub> near 1-yr period peaks before the wavelet-spectrum peak in global temperatures, inconsistent with cooling causing the CO<sub>2</sub> fluctuation. Large eruptions during mid-century (1956 Bezymianny, 1963 Agung, 1964 Kliuchevskoi, 1975 Tolbachik) could be correlated with transient global  $CO_2$  fluctuations near f = 1 cpy in figure 5, but the relative timing is inconsistent. Moreover, a lack of wavelet coherence at f = 1 suggests that the weakly significant coherence in annual-cycle variability in the Fourier analysis is not causally linked to a few correlated transient events.

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